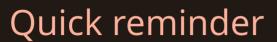
# Multivariate regressions

Lecture 9

Louis SIRUGUE

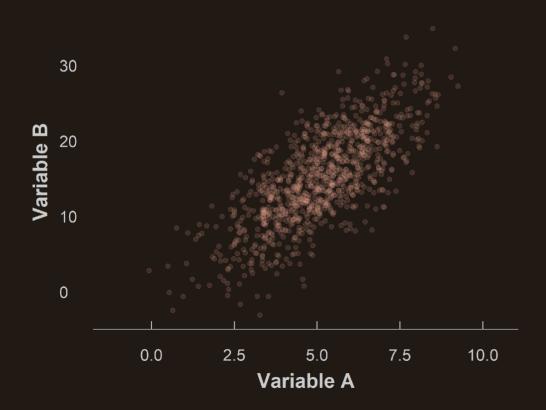
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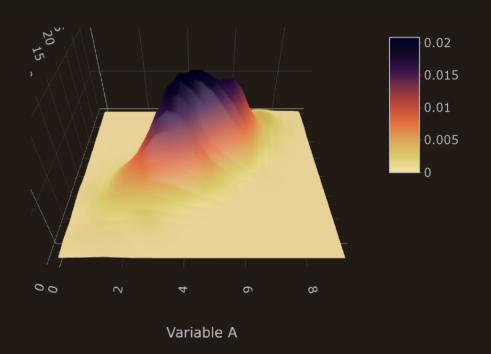


# Ħ

### 1. Joint distribution

The **joint distribution** shows the possible **values** and associated **frequencies** for **two variables** simultaneously







# Quick reminder

### 1. Joint distribution

- → When describing a joint distribution, we're interested in the relationship between the two variables
- The **covariance** quantifies the joint deviation of two variables from their respective mean
  - $\circ$  It can take values from  $-\infty$  to  $\infty$  and depends on the unit of the data

$$ext{Cov}(x,y) = rac{1}{N} \sum_{i=1}^N (x_i - ar{x})(y_i - ar{y}).$$

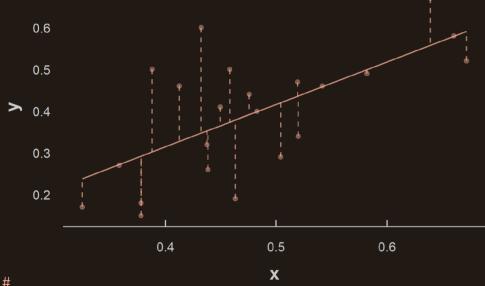
- The **correlation** is the covariance of two variables divided by the product of their standard deviation
  - $\circ~$  It can take values from -1 to 1 and is independent from the unit of the data

$$\operatorname{Corr}(x,y) = rac{\operatorname{Cov}(x,y)}{\operatorname{SD}(x) imes \operatorname{SD}(y)}$$



# Quick reminder

### 2. Regression



• This can be expressed with the **regression** equation:

$$y_i = \hat{lpha} + \hat{eta} x_i + \hat{arepsilon_i}$$

• Where  $\hat{\alpha}$  is the **intercept** and  $\hat{\beta}$  the **slope** of the **line**  $\hat{y_i} = \hat{\alpha} + \hat{\beta}x_i$ , and  $\hat{\varepsilon_i}$  the **distances** between the points and the line

$$\hat{eta} = rac{ ext{Cov}(x_i, y_i)}{ ext{Var}(x_i)}$$

$$\hat{lpha}=ar{y}-\hat{eta} imesar{x}$$

•  $\hat{lpha}$  and  $\hat{eta}$  minimize  $\hat{arepsilon_i}$ 

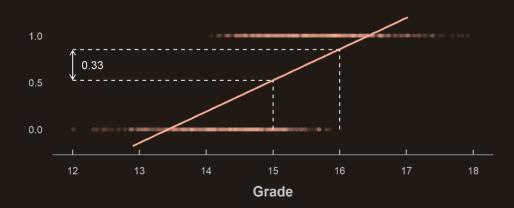




### 3. Binary variables

### Binary **dependent** variables

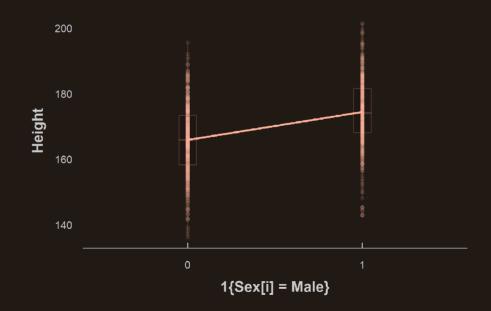
- The **fitted values** can be viewed as **probabilities** 
  - $\circ$   $\hat{eta}$  is the expected increase in the probability that y=1 for a one unit increase in x



• We call that a **Linear Probability Model** 

### Binary **independent** variables

- The x variable should be viewed as a **dummy 0/1** 
  - $\circ \;\; \hat{eta}$  is the difference between the average y for the group x=1 and the group x=0



# Warm up practice

- 1) Open the asec.csv data containing sex, race, weekly work hours, and annual earnings (\$)
- 2) Regress the earnings variable on the sex variable
- 3) Check that the slope coefficient is equal to the difference between male and female average earnings

You've got 10 minutes!

### Solution

1) Open the asec.csv data containing sex, race, weekly work hours, and annual earnings (\$)

```
asec <- read.csv("asec.csv")
```

2) Regress the earnings variable on the sex variable

### Solution

3) Check that the slope coefficient is equal to the difference between male and female average earnings

```
asec %>%

# Group the data by sex
group_by(Sex) %>%

# Summarise mean earnings -> 2x2 dataset
summarise(Mean = mean(Earnings)) %>%

# Put means in columns instead of rows -> 1x2 dataset
pivot_wider(names_from = Sex, values_from = Mean) %>%

# Compute the difference in means
mutate(Difference = Male - Female)
```

```
## # A tibble: 1 x 3
## Female Male Difference
## <dbl> <dbl> <dbl>
## 1 50915. 72527. 21612.
```





- 1.1. Continuous variables
- 1.2. Discrete variables

### 2. Control variables

- 2.1. Motivation
- 2.2. Discrete controls
- 2.3. Continuous controls

### 3. Interactions

- 3.1. Motivation
- 3.2. Discrete interactions
- 3.3. Continuous interactions

### 4. Wrap up!



# Today: Multivariate regressions



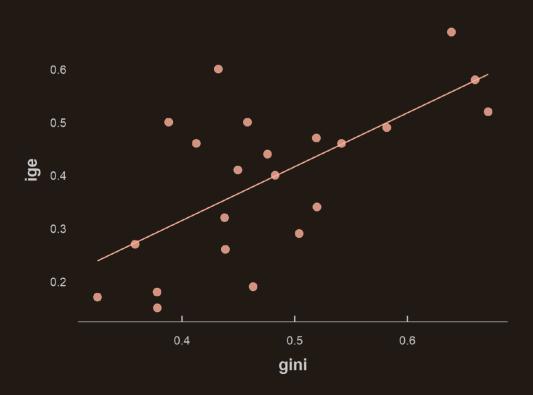
### 1. Adding variables

- 1.1. Continuous variables
- 1.2. Discrete variables

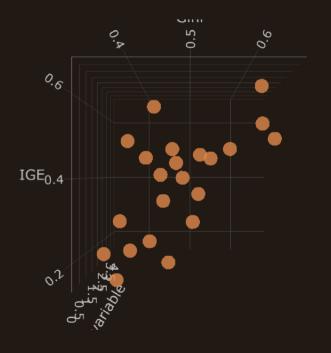


### 1.1. Continuous variables

• So far we focused on two-variable relationships

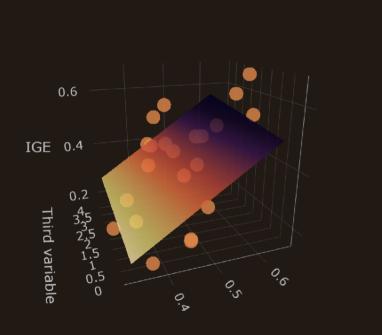


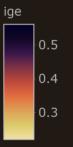
• What about three variable? (pivot the plot)





### 1.1. Continuous variables





- In this case we must fit a **plane** 
  - It is characterized by **3 parameters**
  - And can be expressed as:

$$y_i = \hat{lpha} + \hat{eta_1} x_{1,i} + \hat{eta_2} x_{2,i} + \hat{arepsilon_i}$$

- $\hat{\alpha}$  is still the **intercept** 
  - $\circ~$  The value of  $\hat{y}$  (height) when  $x_1=x_2=0$
- And now there are **2 slopes** 
  - $\circ \;\; \hat{eta_1}$  along the  $x_1$  axis and  $\hat{eta_2}$  along the  $x_2$  axis



#### 1.1. Continuous variables

- The **same** applies with **more than 2** independent variables
  - $\circ$  We would fit a **hyperplane** with as many dimension as x variables
  - $\circ$  We would obtain one intercept and one slope per x variables

$$y_i = \hat{lpha} + \hat{eta_1} x_{1,i} + \hat{eta_2} x_{2,i} {+} \ldots {+} \hat{eta_k} x_{k,i} + \hat{arepsilon_i}$$

- We can estimate the parameters of these hyperplanes in **Im()** 
  - Additional variables must be introduced after a + sign

```
lm(ige ~ gini + third_variable, ggcurve)
```

```
##
## Call:
## lm(formula = ige ~ gini + third_variable, data = ggcurve)
##
## Coefficients:
## (Intercept) gini third_variable
## -0.09536 0.98153 0.01122
```



#### 1.2. Discrete variables

- **So far** we've been working with **binary** categorical variables:
  - Accepted vs. Rejected, Male vs. Female
  - But what about discrete variables with **more than two categories?**
- Take for instance the race variable:

```
asec %>%
  group_by(Race) %>%
  tally()
```

```
## # A tibble: 3 x 2
## Race n
## <chr> <int>
## 1 Black 6835
## 2 Other 6950
## 3 White 50551
```

How can we use this variable as an independent variable in our regression framework?



#### 1.2. Discrete variables

- Remember how we converted our **2-category** variable into **1 dummy** variable
  - We can convert an **n-category** variable into **n-1 dummy** variables

Sex	Male	Race	Black	Other
Female	0	White	0	0
Female	0	White	0	0
Female	0	Black	1	0
Male	1	Black	1	0
Male	1	Other	0	1
Male	1	Other	0	1

### → But why do we omit one category every time?

- Because it would be redundant
- We only need 2 dummies for 3 groups:
  - White: Black = 0 & Other = 0
  - **Black:** Black = **1** & Other = **0**
  - **Other:** Black = **0** & Other = **1**

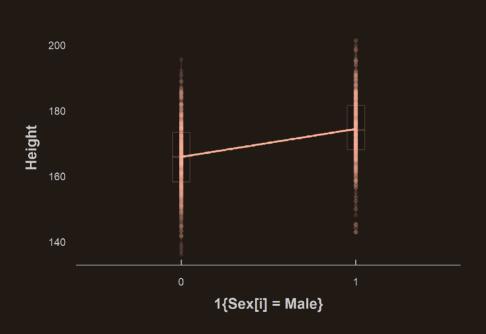
 $\hat{lpha}$  is the expected  $\hat{y}$  when  $x_k=0\ orall k$ 

- Thus is does the job for the omitted groups!
- This group is called the **reference group**
- $\circ$   $\hat{eta}_k$  are interpreted **relative** to that group

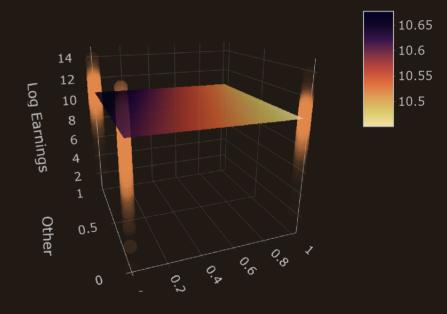


### 1.2. Discrete variables

### 2-category variable



### 3-category variable





#### 1.2. Discrete variables

• This **plane** can be expressed as:

$$ext{Earnings}_i = \hat{lpha} + \hat{eta_1} 1 \{ ext{Race}_i = ext{Other} \} + \hat{eta_2} 1 \{ ext{Race}_i = ext{White} \} + \hat{arepsilon_i}$$

• And the **average** incomes for each group equal:

```
egin{aligned} &\circ & 	ext{Black: } \hat{lpha} + 0\hat{eta}_1 + 0\hat{eta}_2 = \hat{lpha} \ &\circ & 	ext{Other: } \hat{lpha} + 1\hat{eta}_1 + 0\hat{eta}_2 = \hat{lpha} + \hat{eta}_1 \ &\circ & 	ext{White: } \hat{lpha} + 0\hat{eta}_1 + 1\hat{eta}_2 = \hat{lpha} + \hat{eta}_2 \end{aligned}
```

```
##
## Call:
## lm(formula = Earnings ~ Race, data = asec)
##
## Coefficients:
## (Intercept) RaceOther RaceWhite
## 50577 17477 12303
```

Average by group				
Race	Mean earnings			
Black	50577.49			
Other	68054.63			
White	62880.49			



#### 1.2. Discrete variables

- By **default**, lm() sorts categories by **alphabetical** order
  - So every coefficient should be **interpreted relative** to the group which is first alphabetically
- But usually this is **not** the most **intuitive** 
  - You may want everything to be relative to the majority group
  - Or to any group that has reasons to be the **reference**
- The **relevel()** function allows you to **change the reference** category
  - But it works **only on factor** variables

```
##
## Call:
## lm(formula = Earnings ~ Race_fct, data = asec)
##
## Coefficients:
## (Intercept) Race_fctBlack Race_fctOther
## 62880 -12303 5174
```



#### 1.2. Discrete variables

- The **factor class** is made for variables whose values **indicate** different **groups** 
  - Values are just arbitrary group classifiers

```
individuals <- as.factor(c(1, 2, 3, 4, 5))
individuals[1]</pre>
```

```
## [1] 1
## Levels: 1 2 3 4 5
```

- With factors, R understands that the different values do not mean anything
  - And applying standard operations to factors does not make sense

```
individuals * 2

## Warning in Ops.factor(individuals, 2): '*' not meaningful for factors
## [1] NA NA NA NA
```



#### 1.2. Discrete variables

• What you can also do is **create the dummies yourself:** 

```
asec <- asec %>%
  mutate(Black = as.numeric(Race == "Black"),
          Other = as.numeric(Race == "Other"))
lm(Earnings ~ Black + Other, asec)
##
## Call:
## lm(formula = Earnings ~ Black + Other, data = asec)
##
  Coefficients:
                      Black
   (Intercept)
                                   0ther
##
         62880
                     -12303
                                     5174
```

→ This might be the **safest** option



#### 1.2. Discrete variables

• But a **categorical** variable must **not** be introduced **as numeric** in lm()

```
lm(Earnings ~ num_cat, asec)
```

```
##
## Call:
## lm(formula = Earnings ~ num_cat, data = asec)
##
## Coefficients:
## (Intercept) num_cat
## 62093.8 119.6
```

→ lm() used our **categorical** variable as a **continuous** variable



#### 1.2. Discrete variables

• Use the **factor** class

```
asec <- asec %>%
  mutate(fac_cat = as.factor(num_cat))

lm(Earnings ~ fac_cat, asec)

##
## Call:
## lm(formula = Earnings ~ fac_cat, data = asec)
##
## Coefficients:
## (Intercept) fac_cat1 fac_cat2
## 62880 -12303 5174
```

→ Converting all your categorical variables into factors is also a safe option

## Overview



### 1. Adding variables ✓

- 1.1. Continuous variables
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### 2. Control variables

- 2.1. Motivation
- 2.2. Discrete controls
- 2.3. Continuous controls

### 3. Interactions

- 3.1. Motivation
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# Overview



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### 2.1. Motivation

- But **why** would we include **additional variables** in our regressions?
  - The main reason is to **control** for potential **confounders**
- Consider estimating the **relationship** between **income** and exposure air **pollution** in the Paris region

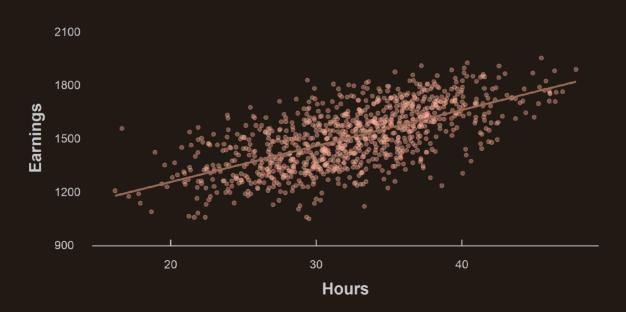
$$\text{Pollution}_i = \hat{lpha_1} + \hat{eta_1} \text{Income}_i + \hat{arepsilon_i}$$

- ullet You would probably expect that  $\hat{eta}_1 < 0$ 
  - Meaning that **higher income** earners live in **less polluted** areas
  - But the closer from Paris the higher the rents and the closer the ring-road
  - $\circ$  This phenomenon might counteract this effect and pull  $\hat{eta}_1$  towards 0
- But how to **remove** the **impact** that **distance** from Paris has on the relationship?
  - **Including it** in the regression would make the corresponding coefficient **absorb the confounding effect**
  - In that case we would call distance a *control* variable

$$ext{Pollution}_i = \hat{lpha_2} + \hat{eta_2} ext{Income}_i + \hat{eta_3} ext{Distance}_i + \hat{\epsilon_i}$$

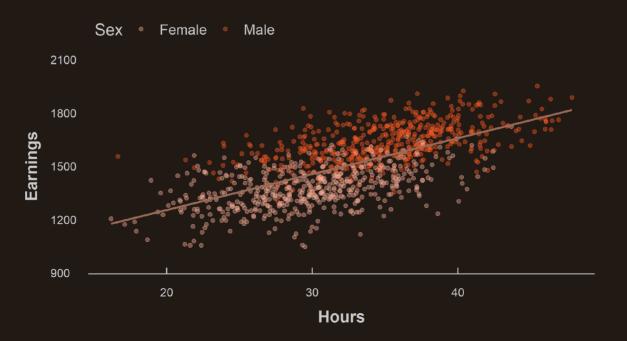


- The most **common control** variable is probably **sex/gender** 
  - It may play a role in the **relationship** between **earnings** and **hours worked** for instance
  - The fact that **women** work **part time** more often and **earn less** contribute to the relationship
  - Just like distance did in the previous example



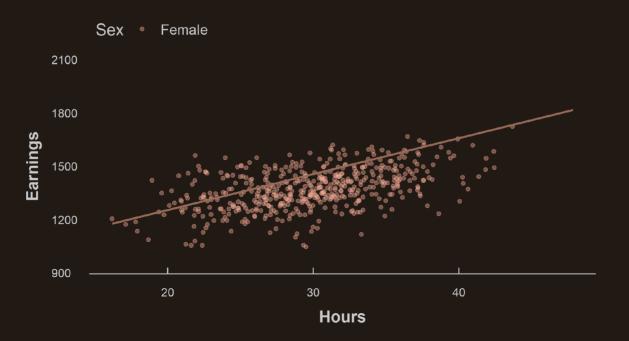


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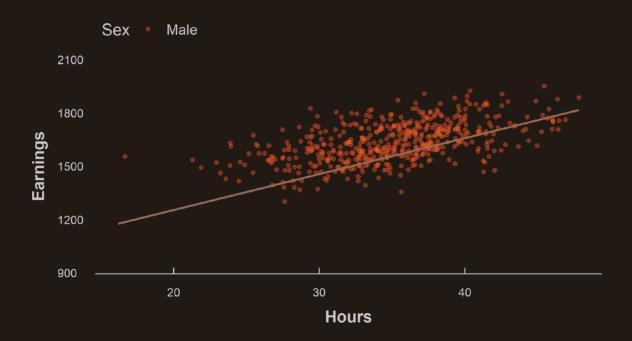


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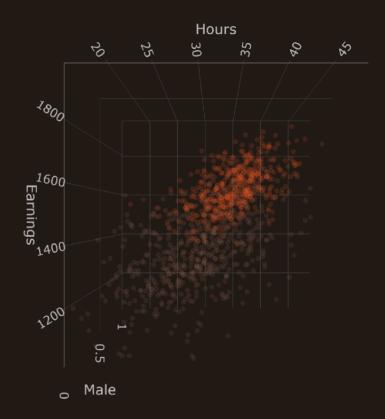


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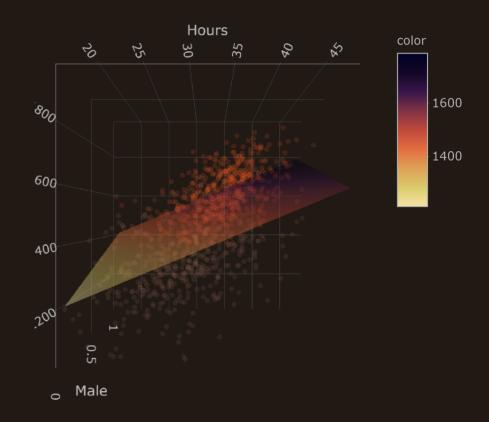


- → The **relationship** is indeed **inflated** by the sex variable
  - Because being a **male** is positively **correlated** with **both** x **and** y
  - Controlling for sex would solve that problem by absorbing this effect
  - Controlling for a discrete variable amounts to allow one intercept per category
  - Giving **two parallel fitted lines** which are the intersections of the plane and the scatterplots





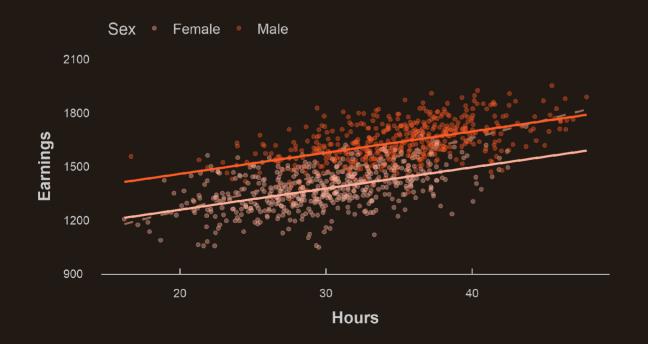
- → The **relationship** is indeed **inflated** by the sex variable
  - Because being a **male** is positively **correlated** with **both** x **and** y
  - Controlling for sex would solve that problem by absorbing this effect
  - Controlling for a **discrete** variable amounts to allow **one intercept per category**
  - Giving **two parallel fitted lines** which are the intersections of the plane and the scatterplots





### 2.2. Discrete

$$ext{Earnings}_i = \hat{lpha} + \hat{eta_1} ext{Hours}_i + \hat{eta_2} 1 \{ ext{Sex}_i = ext{Male} \} + \hat{arepsilon_i}$$



### **Graphical counterpart**

 $\hat{lpha}$ : Intercept of the reference group

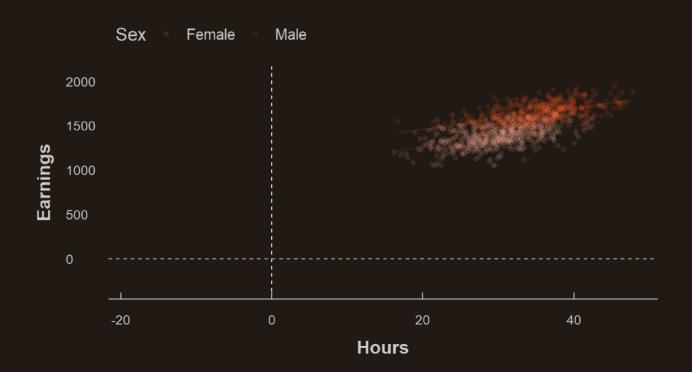
 $\hat{eta_1}$ : Common slope

 $\hat{eta}_2$ : Gap between the two lines

 $\hat{lpha}+\hat{eta_2}$ : Intercept of the other group

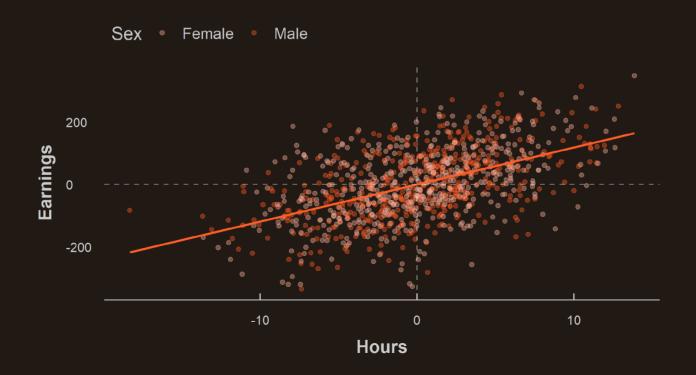


- We can **obtain** this common **slope** by:
  - 1. **Demeaning** earnings and hours by group
  - 2. **Regressing** the demeaned earnings on the hours





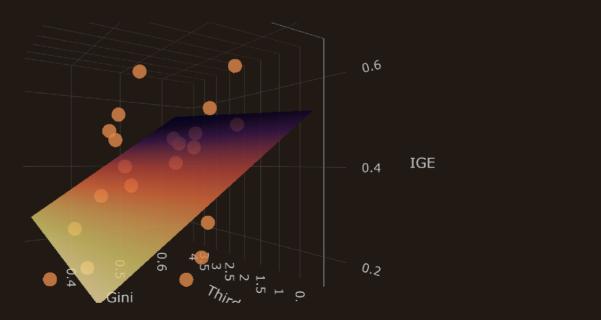
- Note that once we **control** for third variable
  - 1. As we move along the x axis, this **third variable remains constant**
  - 2. Here, as the number of **hours increases** the probability to be a **male does not** increase anymore





#### 2.3. Continuous

- The **same** idea apply when we control for **continuous** variables
  - Including it in the regression allows to **account for another dimension**
  - $\circ$  Such that when x moves this variable **remains constant**
  - $\circ$  This **nets out** the relationship between x and y from the potential **confounding effect** of this variable
  - This is why we call it *controlling* for something



ige

### **Practice**



1) Using the asec data, regress (yearly) earnings on (weekly) hours worked

2) Regress earnings on hours worked controlling for sex

3) Interpret the difference between the results from 1) and 2)

You've got 8 minutes!

# Solution

1) Using the asec data, regress (yearly) earnings on (weekly) hours worked

```
lm(Earnings ~ Hours, asec)$coefficients

## (Intercept) Hours
## -20038.85 2077.79
```

2) Regress earnings on hours worked controlling for sex

```
lm(Earnings ~ Hours + Sex, asec)$coefficients

## (Intercept) Hours SexMale
## -22296.150 1953.829 13794.385
```

# Solution

- 3) Interpret the difference between the results from 1) and 2)
  - The **slope** is still positive **less steep** 
    - In the first regression as the number of hours increases the probability to be a male does increase as well
    - Because males tend to earn more this contributes to the positive relationship between Hours and Earnings
    - In the second regression, controlling for sex allows to maintain the probability to be a male constant along the hour axis to remove this effect

# Overview



# 1. Adding variables ✓

- 1.1. Continuous variables
- 1.2. Discrete variables

### 2. Control variables ✓

- 2.1. Motivation
- 2.2. Discrete controls
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### 3.1. Motivation

- Now we know how to **remove** the **confounding effect** of a third variable by **controlling** for it
  - But what if the main **relationship varies** depending on the value of the **third variable?**
- Let's get back to the previous example

$$ext{Pollution}_i = \hat{lpha} + \hat{eta}_1 ext{Income}_i + \hat{eta}_2 ext{Distance}_i + \hat{\epsilon}_i$$

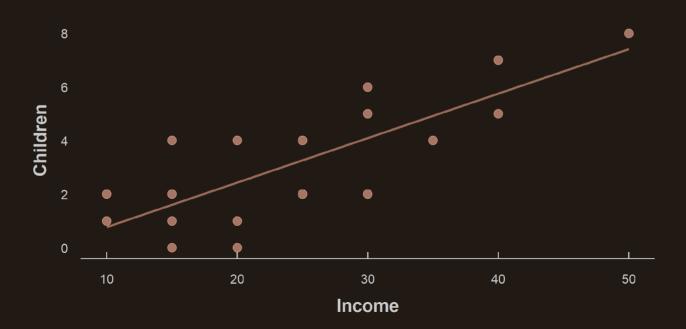
- ullet The **equation imposes** that the **effect** of income on pollution is **constant:**  $eta_2$ 
  - But what if the relationship was actually not the same close to Paris than further away?
  - Maybe that the closer from Paris the larger the effect (higher segregation, ...)
- But how to **capture how the relationship** between income and pollution varies with distance?
  - We should allow for it in the equation!
  - o By adding a term that depends both on income and distance
  - What we use is their **product**, and we call that an **interaction**

$$ext{Pollution}_i = \hat{lpha_2} + \hat{eta_3} ext{Income}_i + \hat{eta_4} ext{Distance}_i + \hat{eta_5} ( ext{Distance}_i imes ext{Income}_i) + \hat{\epsilon_i}$$



# 3.2. Discrete

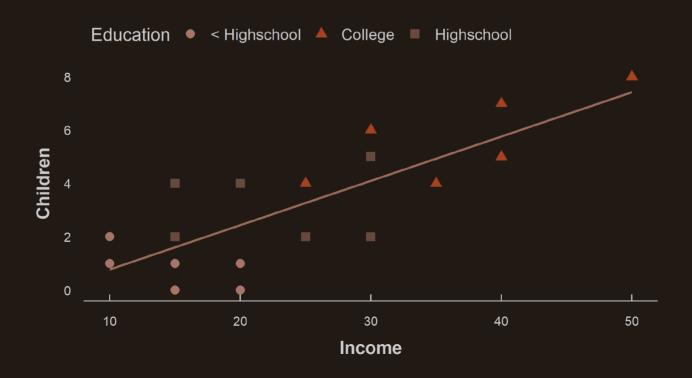
• Take for instance the following **relationship** between **household income** and the **number of children** 





# 3.2. Discrete

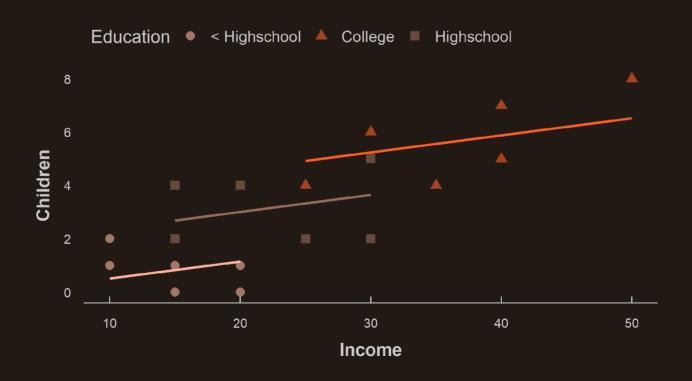
- Take for instance the following **relationship** between **household income** and the **number of children** 
  - The level of **education** seems to **play a role** in the relationship





### 3.2. Discrete

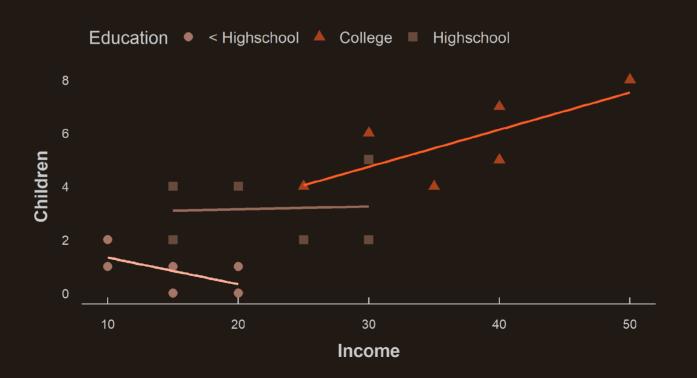
- Take for instance the following **relationship** between **household income** and the **number of children** 
  - o The level of education seems to play a role in the relationship
  - But simply **controlling** for education does **not** seem **sufficient**





### 3.2. Discrete

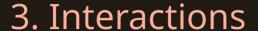
- This is because the **relationship** between income and children **varies with education** 
  - **Interacting** income with education allows to **account for that**
  - Like **controlling** allows for **different intercepts**, **interacting** allows for **different slopes**





### 3.2. Discrete

$$\begin{aligned} \text{Children}_i &= \hat{\alpha_A} + \hat{\beta_A} \text{Income}_i + \\ &\hat{\beta_B} \text{Highschool}_i + \hat{\beta_C} \text{College}_i + \\ &\text{Income}_i \times \left[ \hat{\beta_D} \text{Highschool}_i + \hat{\beta_E} \text{College}_i \right] + \hat{\varepsilon_i} \end{aligned} \qquad \text{Allow for } \neq \text{slopes} \end{aligned}$$





### 3.2. Discrete

→ It is clearly equivalent to regressing children on income separately per education group

$$\begin{aligned} \text{Children}_i &= \hat{\alpha_A} + \hat{\beta_A} \text{Income}_i + \\ \hat{\beta_B} \underbrace{\text{Highschool}_i}_0 + \hat{\beta_C} \underbrace{\text{College}_i}_0 + \\ &\text{Income}_i \times \left[\hat{\beta_D} \underbrace{\text{Highschool}_i}_0 + \hat{\beta_E} \underbrace{\text{College}_i}_0\right] + \hat{\varepsilon_i} \end{aligned} \qquad \text{Allow for } \neq \text{ slopes}$$

< Highschool:  $\operatorname{Children}_i = \hat{lpha_A} + \hat{eta_A} \overline{\operatorname{Income}_i + \hat{arepsilon_i}}$ 



### 3.2. Discrete

$$\begin{aligned} \text{Children}_i &= \hat{\alpha_A} + \hat{\beta_A} \text{Income}_i + \\ \hat{\beta_B} \underbrace{\text{Highschool}_i}_1 + \hat{\beta_C} \underbrace{\text{College}_i}_0 + \\ &\text{Income}_i \times \left[\hat{\beta_D} \underbrace{\text{Highschool}_i}_1 + \hat{\beta_E} \underbrace{\text{College}_i}_0\right] + \hat{\varepsilon_i} \end{aligned} \quad \text{Allow for } \neq \text{ slopes} \end{aligned}$$

**Highschool:** Children
$$_i=(\hat{lpha_A}+\hat{eta_B})+(\hat{eta_A}+\hat{eta_D})\mathrm{Income}_i+\hat{eta_i}$$



### 3.2. Discrete

$$\begin{aligned} \text{Children}_i &= \hat{\alpha_A} + \hat{\beta_A} \text{Income}_i + \\ \hat{\beta_B} \underbrace{\text{Highschool}_i}_0 + \hat{\beta_C} \underbrace{\text{College}_i}_1 + \\ &\text{Income}_i \times \left[\hat{\beta_D} \underbrace{\text{Highschool}_i}_0 + \hat{\beta_E} \underbrace{\text{College}_i}_1 \right] + \hat{\varepsilon_i} \end{aligned} \quad \text{Allow for } \neq \text{ slopes}$$

College: 
$$\hat{\text{Children}}_i = (\hat{\alpha_A} + \hat{eta_C}) + (\hat{eta_A} + \hat{eta_E}) \text{Income}_i + \hat{eta_i}$$



### 3.2. Discrete

$$\begin{split} \text{Children}_i &= \hat{\alpha_A} + \hat{\beta_A} \text{Income}_i + \\ &\qquad \qquad \hat{\beta_B} \text{Highschool}_i + \hat{\beta_C} \text{College}_i + \\ &\qquad \qquad \text{Income}_i \times \left[\hat{\beta_D} \text{Highschool}_i + \hat{\beta_E} \text{College}_i\right] + \hat{\varepsilon_i} \end{split} \qquad \text{Allow for } \neq \text{slopes} \end{split}$$

< Highschool: Children<sub>i</sub> = 
$$\hat{\alpha_A} + \hat{\beta_A} \text{Income}_i + \hat{\varepsilon_i}$$
Highschool: Children<sub>i</sub> =  $(\hat{\alpha_A} + \hat{\beta_B}) + (\hat{\beta_A} + \hat{\beta_D}) \text{Income}_i + \hat{\varepsilon_i}$ 
College: Children<sub>i</sub> =  $(\hat{\alpha_A} + \hat{\beta_C}) + (\hat{\beta_A} + \hat{\beta_E}) \text{Income}_i + \hat{\varepsilon_i}$ 





### 3.3. Continuous

• The **same principle** applies to **continuous variables**:

$$ext{Pollution}_i = \hat{lpha} + \hat{eta}_1 ext{Income}_i + \hat{eta}_2 ext{Distance}_i + \hat{eta}_3 ( ext{Distance}_i imes ext{Income}_i) + \hat{\epsilon_i}$$

• What is the **effect of** a 1-unit increase in **income here?** 

$$\hat{eta_1} + \hat{eta_3} \mathrm{Distance}_i$$

- The **coefficient** associated with the **interaction**,  $\hat{\beta}_3$ , indicates:
  - By how the **effect** of a 1-unit increase in **income** on pollution **varies with distance**
  - $\circ$  When **distance = 0** the effect of income is  $\hat{eta_1}$
  - $\circ~$  For every **additional unit** of distance, the effect of income on pollution **increases by**  $\hat{eta}_3$

→ Don't omit to include your interaction variable as a control in the regression

# Overview



# 1. Adding variables ✓

- 1.1. Continuous variables
- 1.2. Discrete variables

### 2. Control variables ✓

- 2.1. Motivation
- 2.2. Discrete controls
- 2.3. Continuous controls

### 3. Interactions ✓

- 3.1. Motivation
- 3.2. Discrete interactions
- 3.3. Continuous interactions

# 4. Wrap up!



# 4. Wrap up!

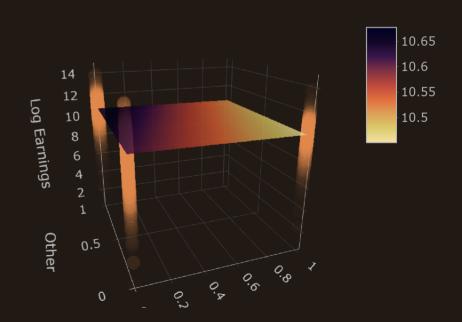
# 1. Multivariate regressions

- Adding a second independent variable in the regression amounts to fitting a plane instead of a line
  - Adding a third variable would fit a hyperplane of dimension 3 and so on

# Adding a continuous variable

# 0.6 0.6 0.4 0.2 Third variable 0.5 0.6 0.6

# Adding a discrete variable



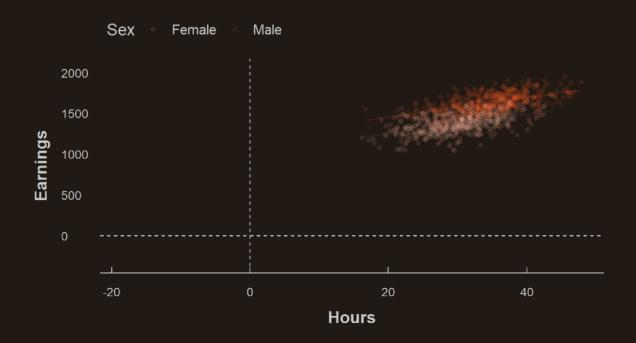


# 4. Wrap up!

# 2. Control variables

- Adding a third variable z removes its potential confounding effect from the relationship between x and y
  - $\circ~$  As we move along the x axis, the **third variable remains constant**

$$\hat{y_i} = \hat{lpha} + \hat{eta_1} x + \hat{eta_2} z + \hat{arepsilon_i}$$





# 4. Wrap up!

### 3. Interactions

• Adding an **interaction** term with z allows to see **how the effect** of x on y varies with z  $\circ$  If z is **discrete**, it amounts to **regressing** y on x **separately** for each z group

$$\hat{y_i} = \hat{lpha} + \hat{eta_1} x + \hat{eta_2} z + \hat{eta_3} (x imes z) + \hat{arepsilon_i}$$

